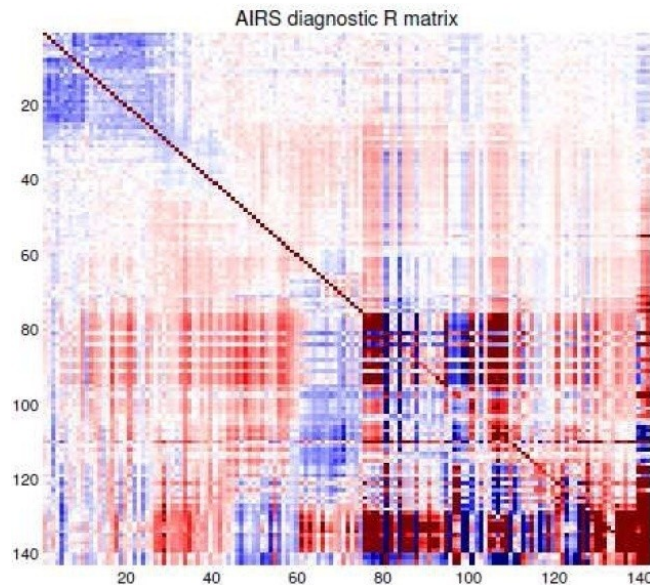


# Diagnosis, Conditioning and Regularization of Error Covariances



**Nancy Nichols\***

Joanne Waller\*, Jemima Tabcart\*, Sarah Dance\*, Amos Lawless\*

# Optimal Bayesian Estimate

**Minimize** with respect to initial state  $\mathbf{x}_0$  :

$$J(\mathbf{x}_0) = \frac{1}{2}(\mathbf{x}_0 - \mathbf{x}_0^b)^T \mathbf{B}^{-1}(\mathbf{x}_0 - \mathbf{x}_0^b) + \frac{1}{2}(\mathcal{H}(\mathbf{x}_0) - \mathbf{y})^T \mathbf{R}^{-1}(\mathcal{H}(\mathbf{x}_0) - \mathbf{y})$$

- Background state  $\mathbf{x}_0^b$
- Observations  $\mathbf{y}$
- Observation operator  $\mathcal{H}$
- Error covariance matrices  $\mathbf{B}$ ,  $\mathbf{R}$

The solution at the **minimum**,  $\mathbf{x}^a$ , is the **analysis**.

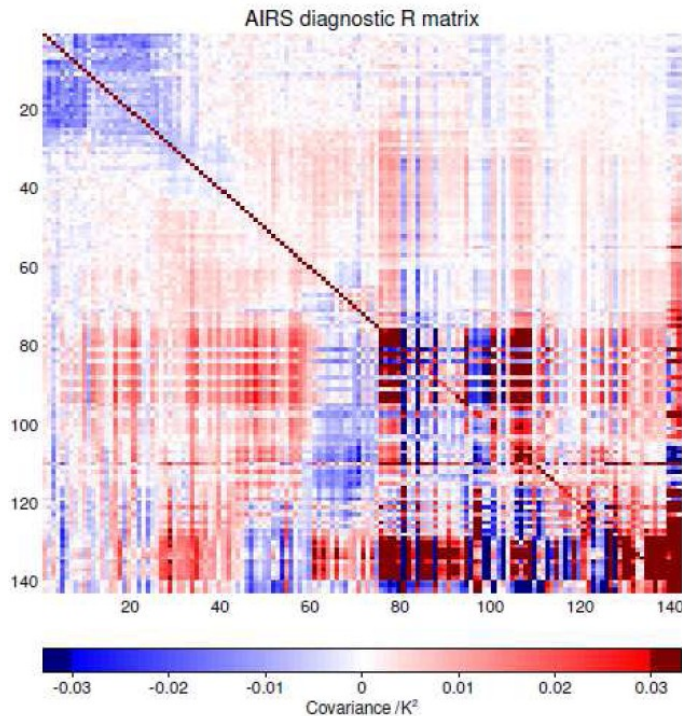


# Outline

- Observation Errors
- Diagnosing Observation Error Covariances
- Incorporating Observation Errors in DA
- Sensitivity of the Analysis
- Regularization
- Conclusions

# 1. Observation Errors

# Observation Errors

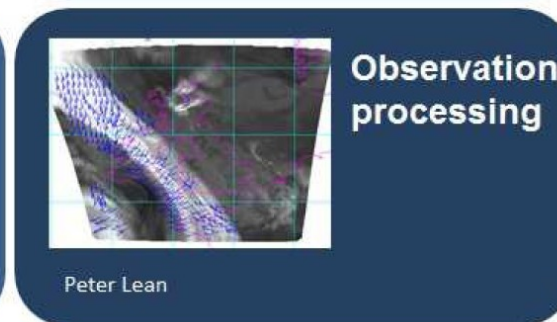
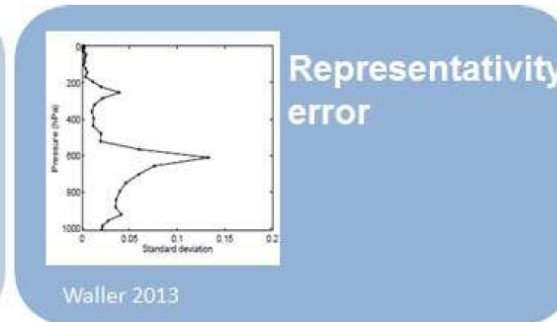


Observation Error Covariance Matrix

- Observation errors have been assumed to be uncorrelated in data assimilation.
- Observation errors in real data are found to be correlated.  
(*Stewart et al, 2009, 2013; Bormann et al, 2010; Waller et al, 2013, 2014a.*)
- Using observation error correlations in data assimilation is shown to improve the state estimate.  
(*Stewart et al, 2008, 2010, 2014; Weston, 2014.*)

# Observation Errors

Four main sources of observation errors, which are **time** and **spatially varying**:



*Waller et al, 2014a; Stewart, 2014; Hodyss & Nichols, 2014*

# Observation Errors

It is important to be able to account for observation error correlations:

- Avoids thinning (high resolution forecasting)
- More information content
- Better analysis accuracy
- Improved forecast skill scores

*Stewart et al, 2008, 2009, 2010, 2013, 2014; Bormann et al, 2010; Waller et al, 2013, 2014a; Weston, 2014*

## 2. Diagnosing Observation Error Covariances



# DBCP Diagnostic (Desroziers et al, 2005)

Let

$$\mathbf{d}_b^o = \mathbf{y} - \mathcal{H}(\mathbf{x}^b),$$
$$\mathbf{d}_a^o = \mathbf{y} - \mathcal{H}(\mathbf{x}^a).$$



# DBCP Diagnostic (Desroziers et al, 2005)

Let

$$\mathbf{d}_b^o = \mathbf{y} - \mathcal{H}(\mathbf{x}^b),$$

$$\mathbf{d}_a^o = \mathbf{y} - \mathcal{H}(\mathbf{x}^a).$$

Then

$$E[\mathbf{d}_a^o \mathbf{d}_b^{oT}] = \tilde{\mathbf{R}}(\mathbf{H}\tilde{\mathbf{B}}\mathbf{H}^T + \tilde{\mathbf{R}})^{-1}(\mathbf{H}\mathbf{B}\mathbf{H}^T + \mathbf{R}) = \mathbf{R}^e,$$



where

$\mathbf{B}$  and  $\mathbf{R}$  are the exact background and observation covariance matrices.

$\tilde{\mathbf{R}}$  and  $\tilde{\mathbf{B}}$  are the assumed observation and background error statistics used in the assimilation.

# DBCP Diagnostic (Desroziers et al, 2005)

Let

$$\mathbf{d}_b^o = \mathbf{y} - \mathcal{H}(\mathbf{x}^b),$$

$$\mathbf{d}_a^o = \mathbf{y} - \mathcal{H}(\mathbf{x}^a).$$

Then

$$E[\mathbf{d}_a^o \mathbf{d}_b^{oT}] = \tilde{\mathbf{R}}(\mathbf{H}\tilde{\mathbf{B}}\mathbf{H}^T + \tilde{\mathbf{R}})^{-1}(\mathbf{H}\mathbf{B}\mathbf{H}^T + \mathbf{R}) = \mathbf{R}^e,$$



If the observation and forecast errors used in the assimilation are exact,  $\tilde{\mathbf{R}} = \mathbf{R}$  and  $\tilde{\mathbf{B}} = \mathbf{B}$ , then

$$E[\mathbf{d}_a^o \mathbf{d}_b^{oT}] = \mathbf{R}$$

# DBCP Diagnostic in Spectral Space

**Analysis** of the diagnostic in spectral space, under some simplifying assumptions, shows that **if** the observation errors are **correlated**, then assuming in the assimilation that the correlation matrix is **diagonal** results in an estimate  $R^e$  with: :

- **underestimated** observation error **variances**;
- **underestimated** observation error **correlation length scales**;

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**Analysis** of the diagnostic in spectral space, under some simplifying assumptions, shows that **if** the observation errors are **correlated**, then assuming in the assimilation that the correlation matrix is **diagonal** results in an estimated  $\mathbf{R}^e$  with:

- **underestimated** observation error **variances**;
- **underestimated** observation error **correlation length scales**.

But a **better estimate** of the observation error **covariance** matrix than an uncorrelated diagonal matrix.

*Waller et al, 2016a*

# Summary: DBCP Diagnostic

The DBCP diagnostic has been successfully **applied** in **operational systems** to determine the observation error covariances for a variety of different observation types: including:

- Doppler radar wind data;
- atmospheric motion vectors;
- remotely sensed satellite data –  
eg SEVIRI, IASI, ARES, CRIS and others

*Stewart et al, 2014; Waller et al, 2016b, 2016c; Cordoba et al, 2016.*

# 3. Incorporating Correlated Observation Errors in Ensemble DA

# ETKF Filter

**Step 1** Use the full **non-linear model** to forecast each ensemble member from  $\mathbf{x}_{n-1}^a$  to  $\mathbf{x}_n^f$ .

**Step 2** Calculate the ensemble **mean**  $\bar{\mathbf{x}}_n^f$  and approximate **covariance** matrix  $\mathbf{B}_n$ .

**Step 3** Using the ensemble mean at time  $t_n$ , calculate the **innovation**  $\mathbf{d}_{b_n}^o$ .

**Step 4** The ensemble mean is updated using

$$\bar{\mathbf{x}}_n^a = \bar{\mathbf{x}}_n^f + \mathbf{K}_n \mathbf{d}_{b_n}^o$$

where the gain  $\mathbf{K}_n = \mathbf{Z}_n \mathbf{H}_n^T \mathbf{R}_n^{-1} \approx \mathbf{B}_n \mathbf{H}_n^T (\mathbf{H}_n \mathbf{B}_n \mathbf{H}_n^T + \mathbf{R}_n)^{-1}$

*Living's et al, 2008*



# Ensemble Filter with Diagnostic

## Procedure:

- Select initial  $\mathbf{R}$
- Run ETKF and store samples of  $\mathbf{d}^b$  and  $\mathbf{d}^a$
- Compute  $E[\mathbf{d}^a \mathbf{d}^{bT}]$
- Symmetrize (and **regularize**) to obtain new estimate for  $\mathbf{R}$
- Repeat steps of ETKF using samples from rolling window of length  $N_s$  to update  $\mathbf{R}$

*Waller et al, 2014a*

# Example:

Use high resolution Kuramoto-Sivashinsky model

Add errors to observations from normal distribution with known SOAR covariance  $\mathbf{R}_t$ .

- Assume incorrect  $\mathbf{R}_1 = \text{diagonal}$  at  $t = 0$ .  
Recover fixed true covariance.
- Allow length scale in true covariance to vary slowly. Recover time-varying true covariance.

# Results – Static $R_t$ :

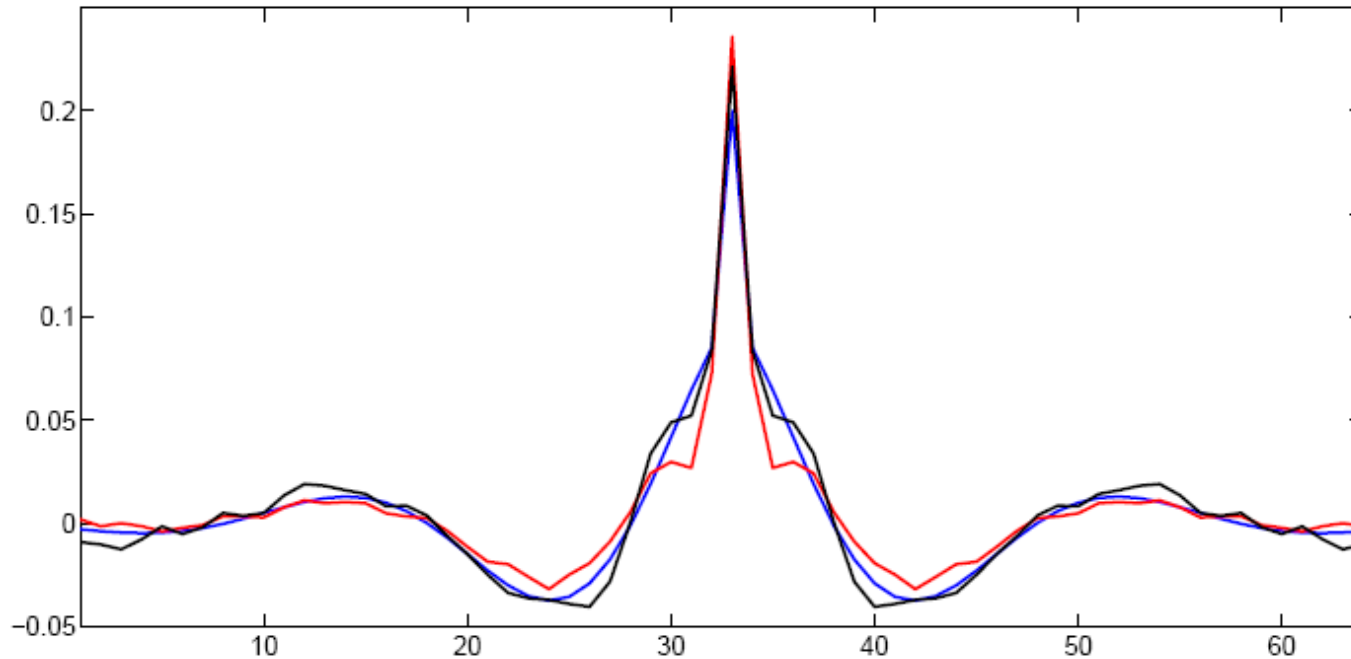
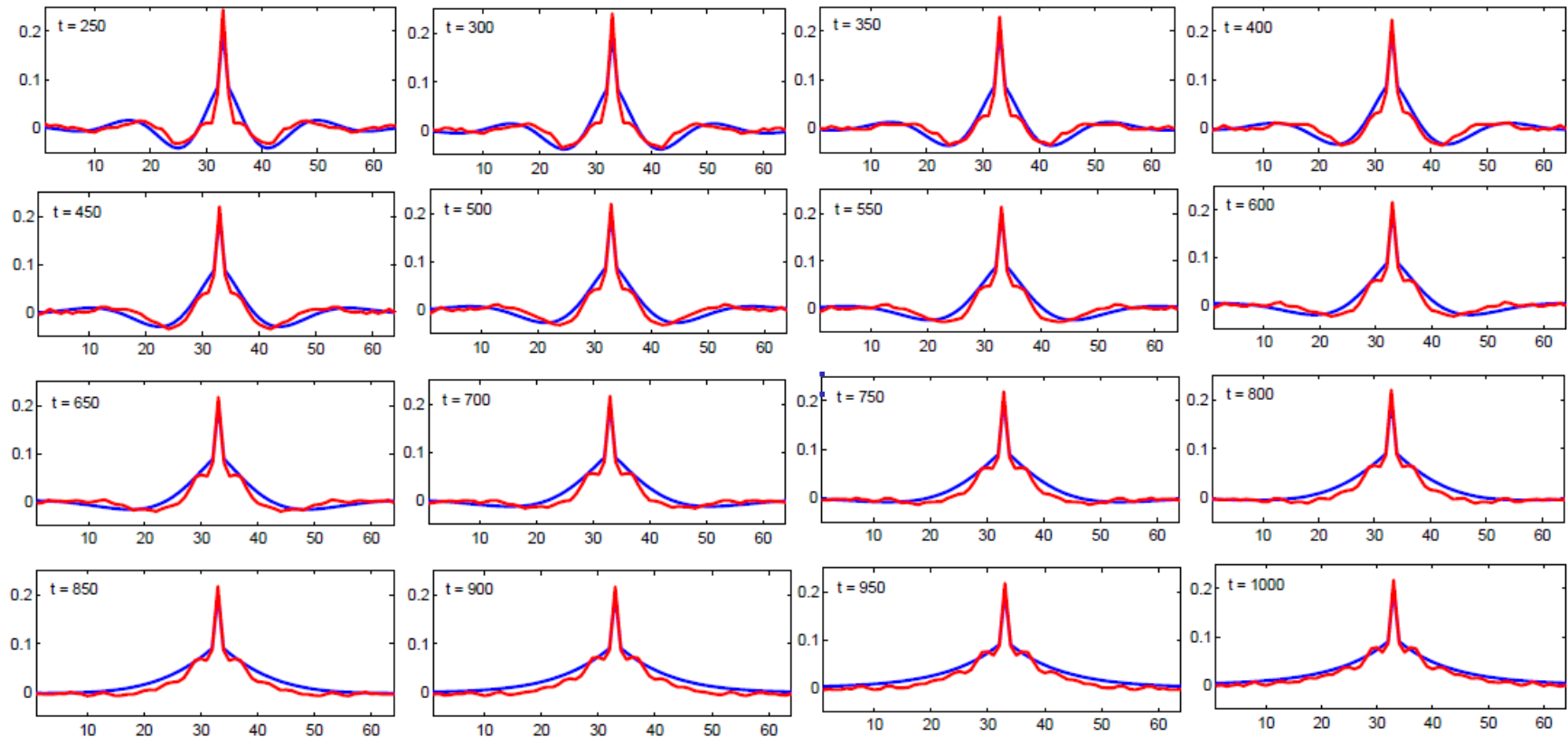


Figure : Rows of the true (blue) and estimated (red, initial estimation, black final estimation) correlation matrices for Experiment 4

# Results – Time Varying $R_t$ :



Rows of the true and estimated correlation matrices

# Results – Analysis Errors:

Time averaged RMSE **analysis** errors:

Static True  $\mathbf{R}_t$

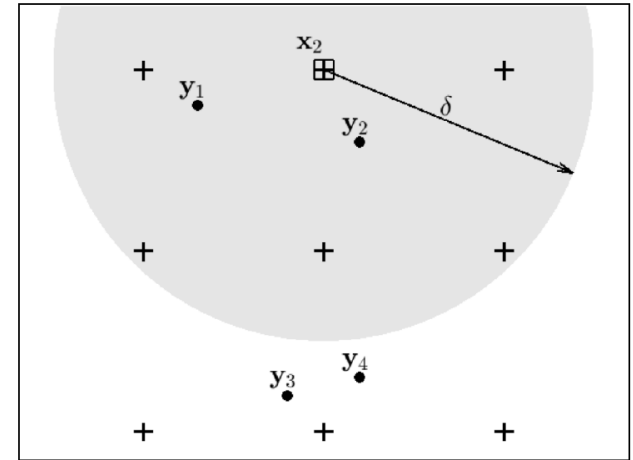
- Experiment: Exact  $\mathbf{R}_t$  0.246
- Experiment:  $\mathbf{R} = \mathbf{R}_t$  fixed 0.275
- Experiment:  $\mathbf{R}$  updated 0.251

Time Varying True  $\mathbf{R}_t$  0.255

**Conclude:** the **analysis** is **improved** by incorporating the estimated observation error covariance in the DA

# Localization and DBCP Diagnostic

Regularization of the matrix  $\mathbf{R}^e$  is needed to ensure stability of the filter. With domain **localization**, states are only updated using observations within a localization radius.

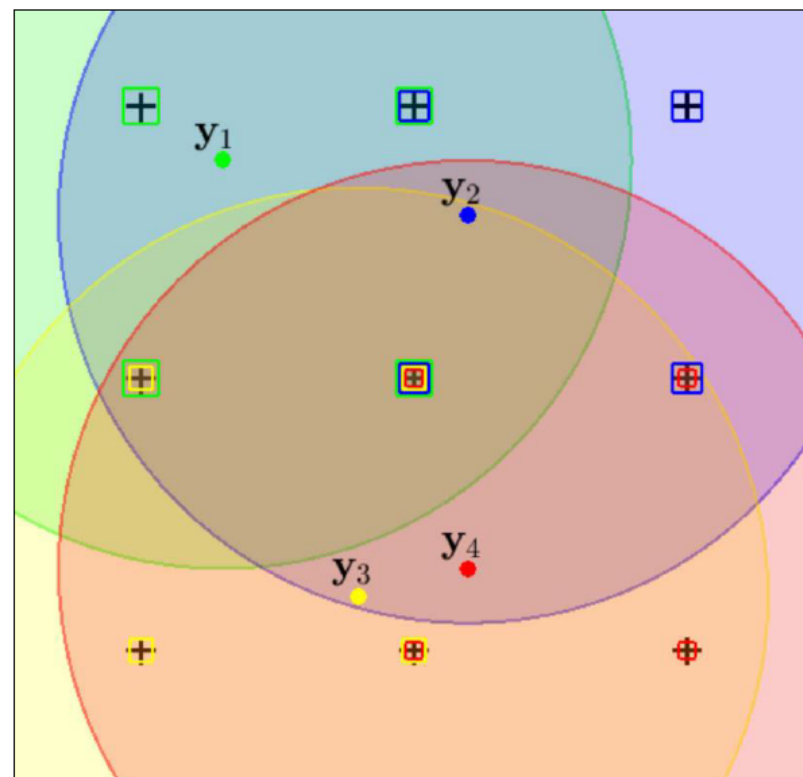


**Caveat:** Computing the DBCP diagnostic using samples from an ensemble filter with domain localization does **not** give the correct values of all the observation error covariances, **even if** all theoretical assumptions hold.

*Waller, Dance & Nichols, 2017*

# Definitions:

- The **domain of dependence** of an observation  $y_i$  is the set of elements of the model state that are used to calculate the model equivalent of  $y_i$ .
- The **region of influence** of an observation is the set of analysis states that are updated in the assimilation using the observation  $y_i$ .



# Definitions:

The DD region is determined by **H**. The RI region is determined by **F** and depends on the radius of localization.

$$\mathbf{H} = \begin{bmatrix} \checkmark & \checkmark & \times & \checkmark & \checkmark & \times & \times & \times & \times \\ \times & \checkmark & \checkmark & \times & \checkmark & \checkmark & \times & \times & \times \\ \times & \times & \times & \checkmark & \checkmark & \times & \checkmark & \checkmark & \times \\ \times & \times & \times & \times & \checkmark & \checkmark & \times & \checkmark & \checkmark \end{bmatrix}$$

$$\mathbf{F} = \begin{bmatrix} \checkmark & \checkmark & \times & \times \\ \checkmark & \checkmark & \times & \times \\ \times & \checkmark & \times & \times \\ \checkmark & \checkmark & \checkmark & \checkmark \\ \checkmark & \checkmark & \checkmark & \checkmark \\ \times & \checkmark & \checkmark & \checkmark \\ \times & \times & \checkmark & \checkmark \\ \times & \times & \checkmark & \checkmark \\ \times & \times & \checkmark & \checkmark \end{bmatrix}$$



# Theorem:

The correlation  $\mathbf{R}_{ij}$  between observations  $y_i$  and  $y_j$  is determined correctly by the DBCP diagnostic **only if** the **domain of dependence** of  $y_i$  lies within the region of influence of observation  $y_j$ .

That is: the  $(i, j)$  element of  $\mathbf{H}(\mathbf{F} - \mathbf{B}\mathbf{H}^T) = 0$ .

*Waller, Dance & Nichols, 2017*

# Summary : DBCP Diagnostic in Ensemble DA

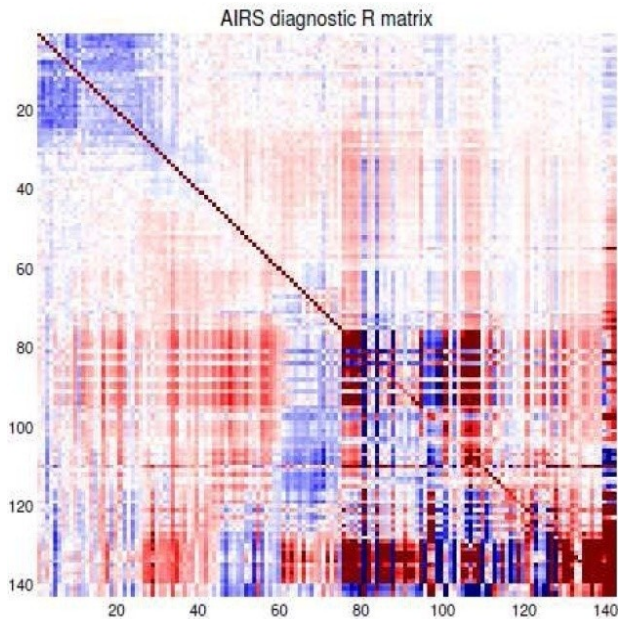
The **DBCP** diagnostic can be used **with care** to estimate the observation error correlation matrix **R** in **ensemble DA**.

In practice the diagnosed matrix **R** may be **ill-conditioned** and may need to be reconditioned.

Accounting for the correlated errors in practice is a **computational challenge**, now being tackled.

# 4. Sensitivity of the Analysis

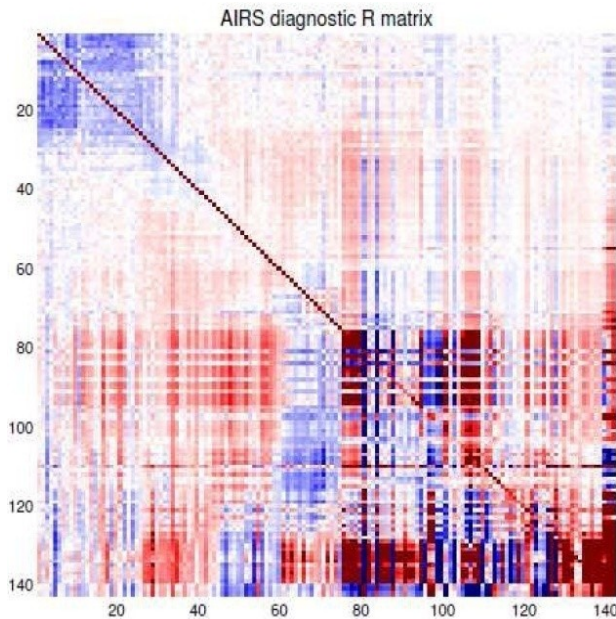
# Problems for DA:



Diagnosed correlation matrices:

- Non-symmetric
- Variances too small
- Not positive-definite
- **Very** ill-conditioned

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Diagnosed correlation matrices:

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- Not positive-definite
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**Aim:** to examine the **sensitivity** of the **analysis** to the **conditioning** of the estimated observation error covariances.

# Sensitivity of the Analysis

Sensitivity of the analysis, is bounded in terms of the condition number of:

$$\mathbf{S} = \mathbf{B}^{-1} + (\mathbf{H})^T \mathbf{R}^{-1} \mathbf{H}$$

where  $\mathbf{B}$  and  $\mathbf{R}$  are covariance matrices with structures that depend on the variances and correlation length scales of the background and observation errors, respectively.

# Sensitivity

We can establish the following **theorem**:

*Let  $\mathbf{B} \in \mathbb{R}^{N \times N}$  and  $\mathbf{R} \in \mathbb{R}^{p \times p}$ , with  $p < N$ , be the background and observation error covariance matrices respectively. Additionally, let  $\mathbf{H} \in \mathbb{R}^{p \times N}$  be the observation operator. Then the following bounds are satisfied by the condition number of the Hessian,  $\mathbf{S} = \mathbf{B}^{-1} + \mathbf{H}^T \mathbf{R}^{-1} \mathbf{H}$ ,*

$$\frac{\kappa(\mathbf{B})}{\left(1 + \frac{\lambda_{\max}(\mathbf{B})}{\lambda_{\min}(\mathbf{R})} \lambda_{\max}(\mathbf{H}\mathbf{H}^T)\right)} \leq \kappa(\mathbf{S}) \leq \left(1 + \frac{\lambda_{\min}(\mathbf{B})}{\lambda_{\min}(\mathbf{R})} \lambda_{\max}(\mathbf{H}\mathbf{H}^T)\right) \kappa(\mathbf{B}).$$

*Haben et al, 2011; Haben 2011; Tabcart, 2016; Tabcart et al, 2018*

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Note: the upper bound **grows** as  $\frac{1}{\lambda_{\min}(\mathbf{R})}$  **grows** and depends also on the observation operator.

*Haben et al, 2011; Haben 2011; Tabcart, 2016; Tabcart et al, 2018*



# Sensitivity

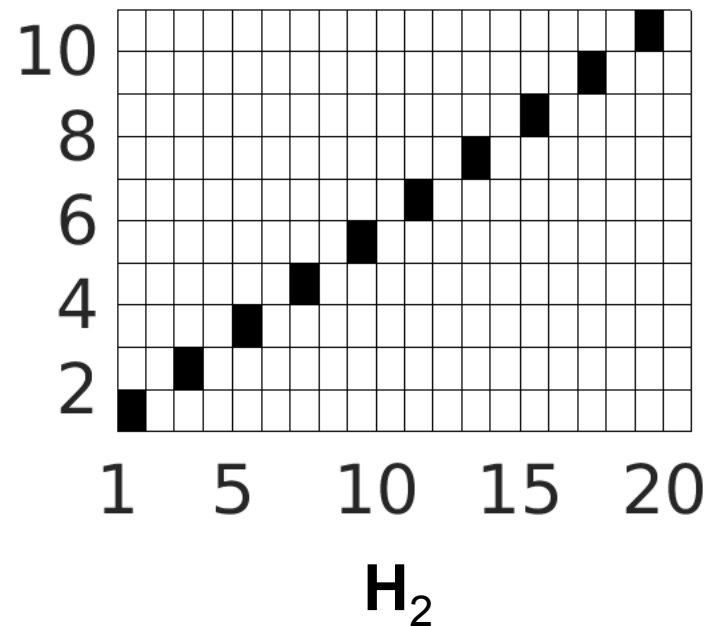
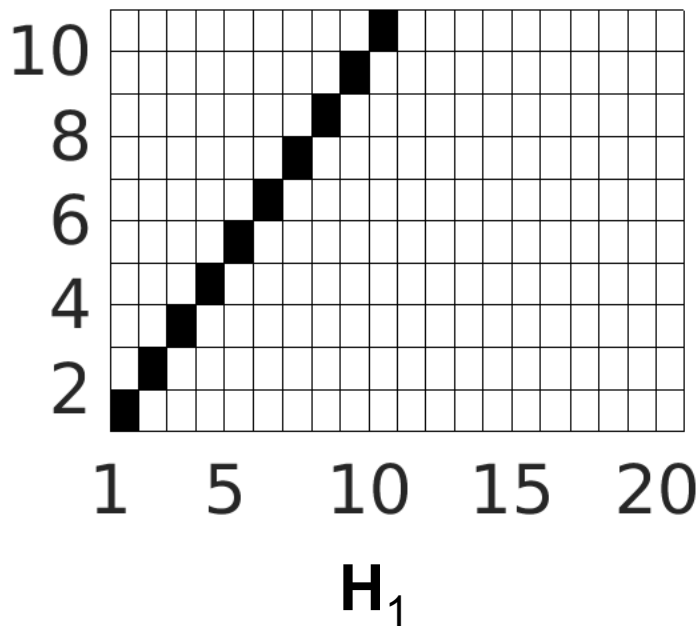
## Key questions:

- What happens when we change the length scales of **R** and **B** - separately? together?
- What affect does the choice of observation operator have?
- How does changing the minimum eigenvalue of **R** affect the conditioning of **S** ? Operationally?

# Example:

We examine how the choice of operator and the length scales in **R** and **B** affect the **sensitivity** of the analysis.

Choice of observation operator:



# Example - $\mathbf{H}_1$ :

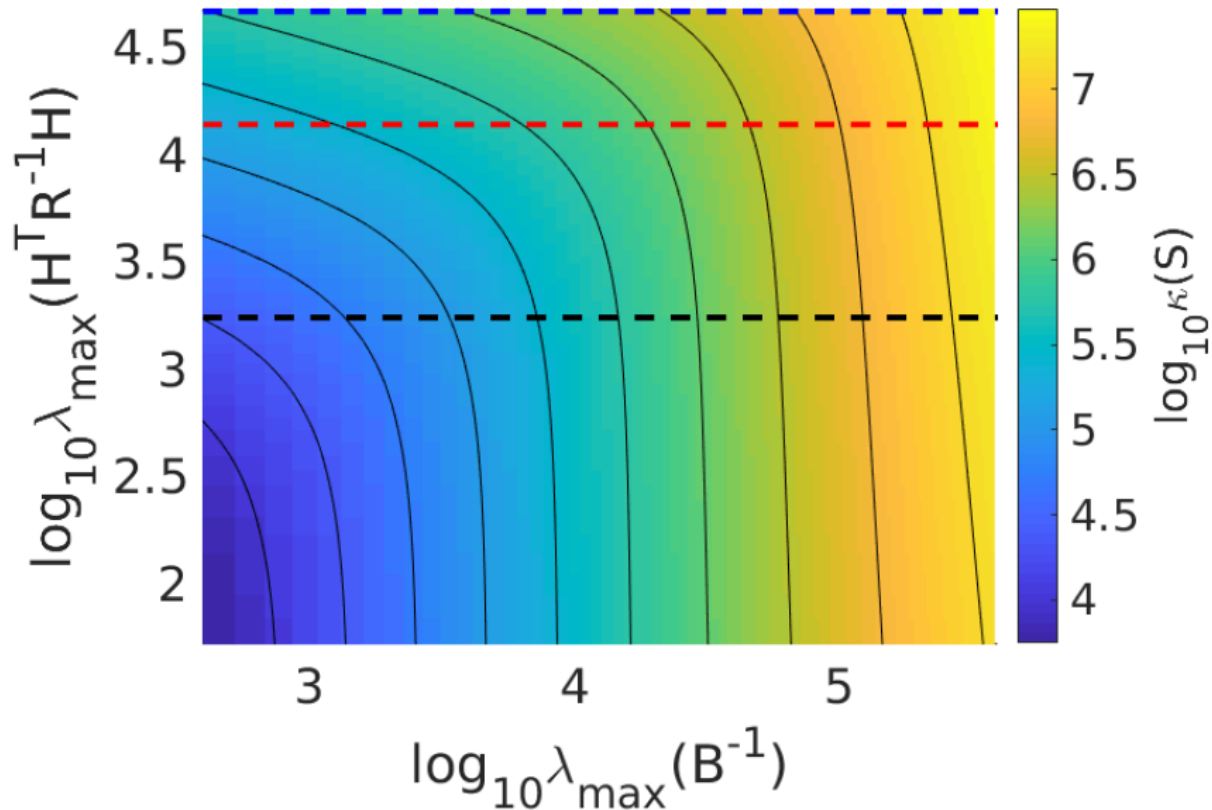


Figure: Plot of changing  $\lambda_{\max}(\mathbf{B}^{-1})$  against  $\lambda_{\max}(\mathbf{H}^T \mathbf{R}^{-1} \mathbf{H})$  with colour denoting the condition number of  $\mathbf{S}$  when first 100 state variables observed.

## Example - $\mathbf{H}_2$ :

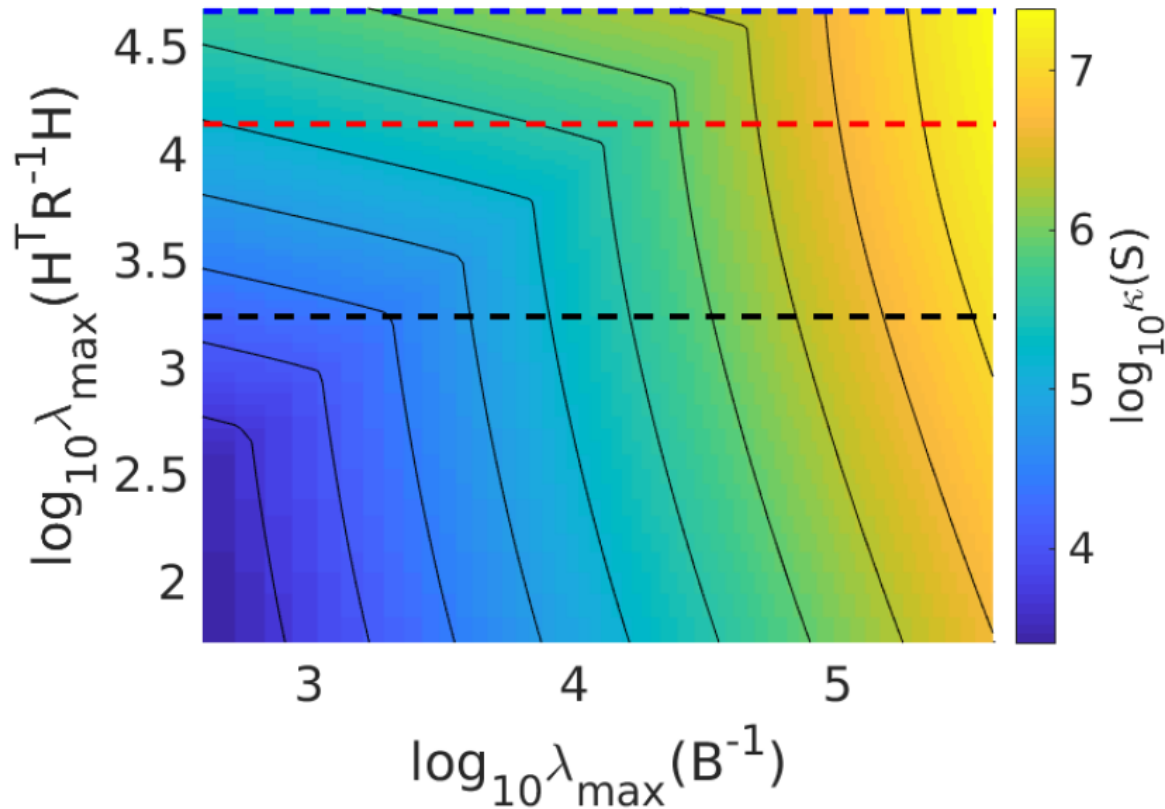


Figure: Plot of changing  $\lambda_{\max}(\mathbf{B}^{-1})$  against  $\lambda_{\max}(\mathbf{H}^T \mathbf{R}^{-1} \mathbf{H})$  with colour denoting the condition number of  $\mathbf{S}$  when alternate state variables observed

# Summary: Conditioning of the Problem

We find that the condition number of **S** **increases** as:

- the observations become **more accurate**
- the observation **length scales increase**
- the prior (background) becomes **less accurate**
- the prior error correlation **length scales increase**
- the observation error covariance becomes **ill-conditioned** - ie when  $\frac{1}{\lambda_{min}(\mathbf{R})}$  becomes **large**

*Haben et al, 2011; Haben 2011; Tabcart, 2016; Tabcart et al, 2018*

# 5. Regularization

# Reconditioning $\mathbf{R}$

To improve the conditioning of  $\mathbf{R}$  (and  $\mathbf{S}$ ) we alter the eigenstructure of  $\mathbf{R}$  so as to obtain a specified condition number for the modified covariance matrix by:

- **Ridge regression (RR)**: add constant to all diagonal elements to achieve given condition number.
- **Eigenvalue modification (ME)**: increase the smallest eigenvalues of  $\mathbf{R}$  to a threshold value to achieve the given condition number, keeping the rest unchanged.

# Theoretical Results:

- Both methods reduce the condition number of  $\mathbf{R}$ .
- Both methods increase all the standard deviations, but ridge regression creates a larger increase than does the eigenvalue modification method.
- Ridge regression decreases the moduli of all the cross-correlations.
- The eigenvalue modification method is equivalent to minimizing the KyFan 1-p (trace) norm of the distance to the nearest covariance matrix with condition number less or equal to a given value  $K_{max}$ .

*Tabcart et al, 2018*



## Example:

Given a covariance matrix, constructed by sampling a SOAR correlation function, with condition number 81121 and fixing the variances to be constant. Recondition using RR and ME.

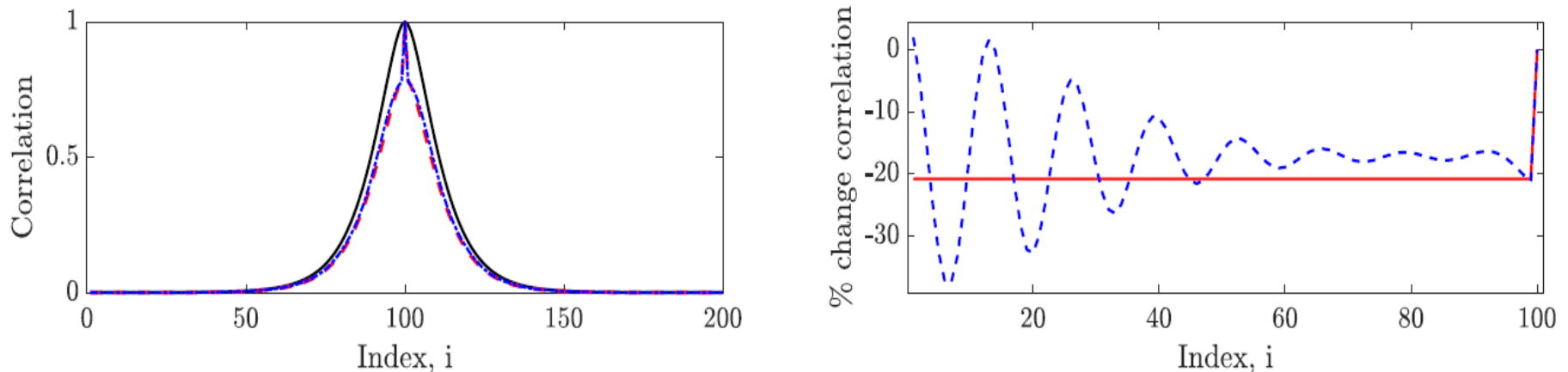
Table 1: Change to standard deviations of the SOAR matrix.

$\kappa_{max}$	$\sigma$	$\sigma_{RR}$	% change RR	$\sigma_{ME}$	% change ME
1000	2.23606	2.26471	+1.28%	2.25439	+0.82%
500	2.23606	2.29340	+2.56%	2.27599	+1.79%
100	2.23606	2.51306	+12.39%	2.45737	+9.90%

# Example:

Given a covariance matrix, constructed by sampling a SOAR correlation function, with condition number 81121 and fixing the variances to be constant. Recondition using RR and ME.

Figure 1: Changes to correlations for  $\kappa_{max} = 100$ .



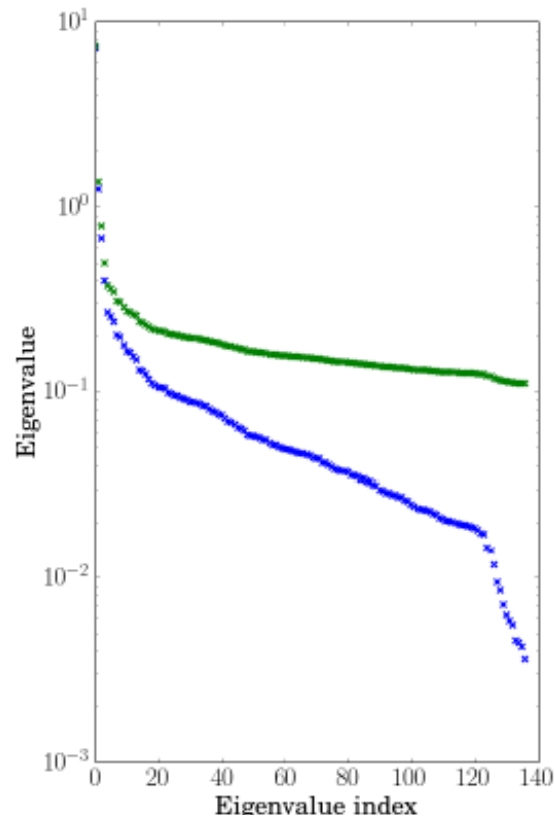
RR = red solid, ME= blue dashed, Original = black solid

# Operational Tests - Met Office

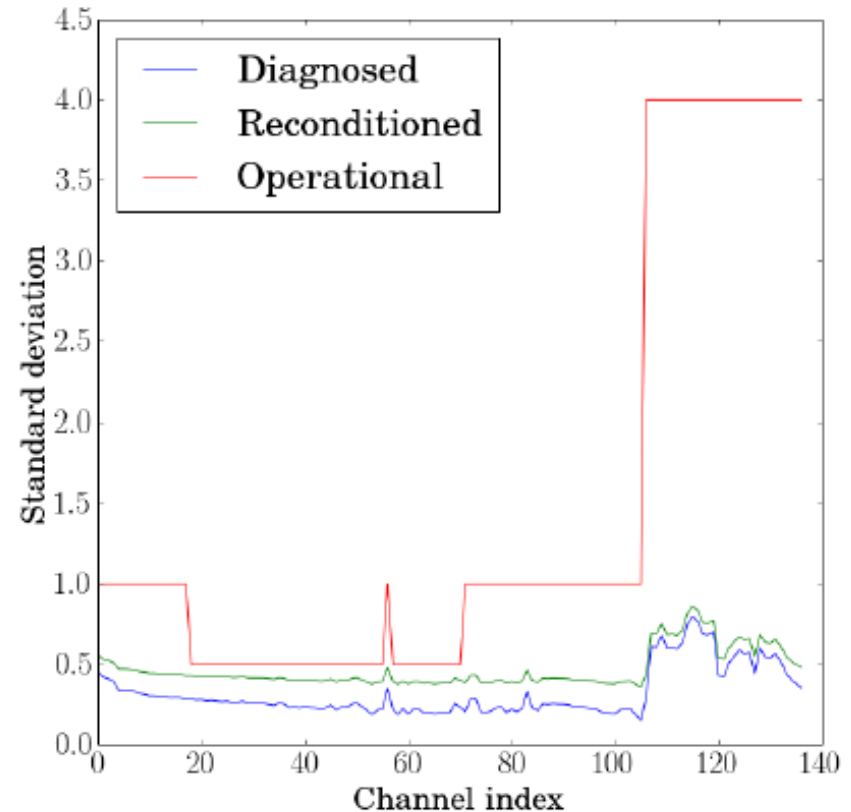
Experiments using the Met Office 1D satellite retrieval system

- Aim to test qualitative conclusions in an operational system.
- Focus on observations from IASI (Infrared Atmospheric Sounding Interferometer) instrument (on MetOp-A satellite). Note the observation operator is non-linear in this case.
- Investigate how changing the minimum eigenvalue of  $\mathbf{R}$  affects the condition number of  $\mathbf{S}$  - we only show results using the ridge regression method.

# Results - 1:



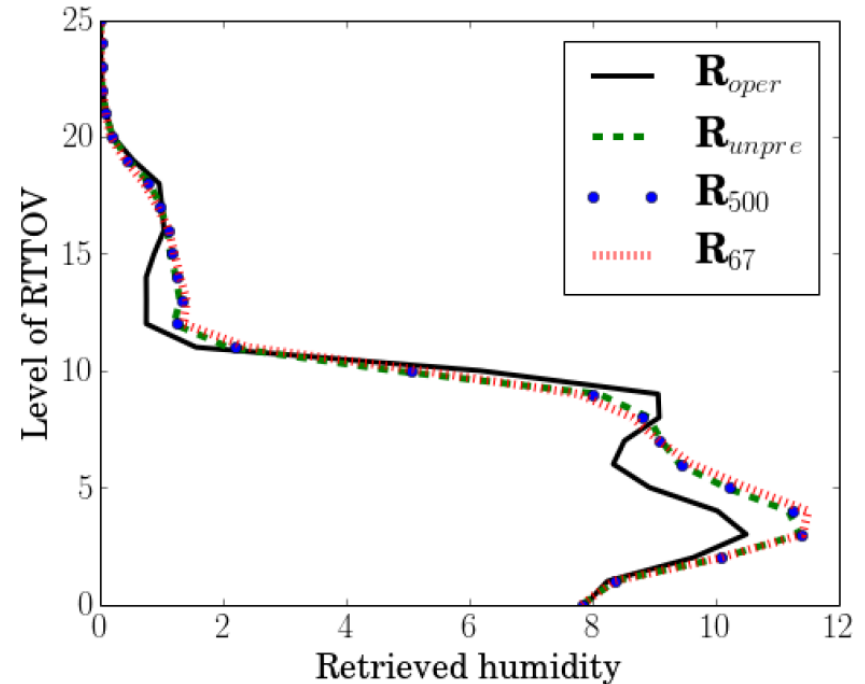
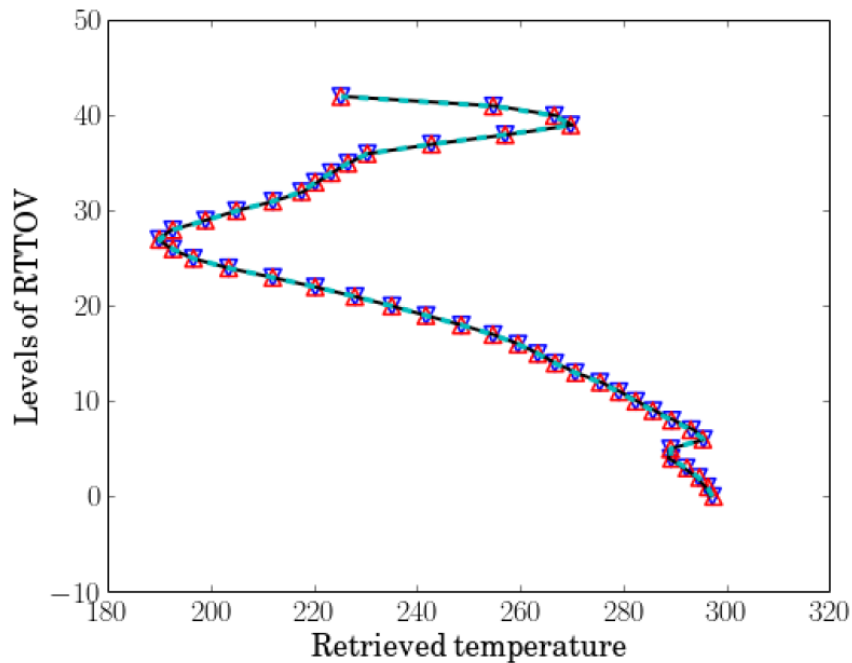
(a) Change in eigenvalues of  $\mathbf{R}^{137}$



(b) Standard deviations for different  $\mathbf{R}^{137}$

The method of reconditioning is that described in Algorithm 1. Figure (a) depicts how the ridge regression method changes the eigenvalues of the covariance matrix  $\mathbf{R}^{137}$ . Eigenvalues of the raw Desroziers  $\mathbf{R}_{unpre}^{137}$  are shown in blue, and those of  $\mathbf{R}_{RC}^{137}$  reconditioned so that  $\kappa(\mathbf{R}_{RC}) = 67$  are shown in green. Figure (b) compares the standard deviations of  $\mathbf{R}_{old}^{137}$  (red),  $\mathbf{R}_{unpre}^{137}$  (blue), and  $\mathbf{R}_{67}^{137}$  (green).

# Results - 2



Shown are the retrieved temperature and humidity profiles for 4 different choices of R:  $R_{oper}$ ,  $R_{unpre}$ ,  $R_{500}$  and  $R_{67}$ .

# Summary: Regularization

- Developed theory on reconditioning of the matrix  $\mathbf{R}$ .
- Theory tested in a twin experiment – showing effect of ridge regression and eigenvalue modification on standard deviations and correlations of the modified covariance matrices.
- Operationally standard deviations of the diagnosed matrices are increased by the reconditioning . The impact on temperature retrievals was minimal, but the impact on humidity retrievals much larger.

*Tabcart, 2016; Tabcart et al, 2018b*

# 6. Conclusions

# Conclusions

**Ensemble DA** allows the statistical estimation of the background and observation covariance matrices from sampled states.

In practice the diagnosed matrices are commonly singular or very **ill-conditioned**. **Regularization** is required to ensure the **stability** of the filter.

A variety of techniques are available, including **localization** and **reconditioning**. A combination of these two approaches have been applied to an 4DEnVar simplified system and shown to be of benefit.

*Smith et al, 2017*



Many more challenges left!



# References

- Bormann N and Bauer P. 2010. Estimates of spatial and interchannel observation-error characteristics for current sounder radiances for numerical weather prediction. I: Methods and application to ATOVS data. *QJ Royal Meteor Soc*, 136:1036–1050.
- Bormann N, Collard A, Bauer P. 2010. Estimates of spatial and interchannel observation-error characteristics for current sounder radiances for numerical weather prediction II: application to AIRS and IASI data. *QJ Royal Meteor Soc* 136: 1051 – 1063.
- Cordoba M, Dance SL, Kelly GA, Nichols NK and Waller JA. 2017. Diagnosing Atmospheric Motion Vector observation errors for an operational high resolution data assimilation system", *Quarterly Journal of the Royal Meteorological Society, Part A*, 143, 333–341.
- Desroziers G, Berre L, Chapnik B, Poli P. 2005. Diagnosis of observation, background and analysis-error 131: 3385 – 3396.
- Desroziers G, Berre L and Chapnik B. 2009. Objective validation of data assimilation systems: diagnosing sub-optimality. In: *Proceedings of ECMWF Workshop on diagnostics of data assimilation system performance*, 15-17 June 2009.
- Hodyss D and Nichols NK. 2015. Errors of representation: basic understanding, *Tellus A*, 67, 24822 (17 pp)
- Haben SA, Lawless, AS and Nichols NK. 2011. Conditioning of incremental variational data assimilation, with application to the Met Office system, *Tellus*, **63A**, 782 – 792.
- Haben SA. 2011 Conditioning and Preconditioning of the Minimisation Problem in Variational Data Assimilation, PhD thesis, Dept of Mathematics & Statistics, University of Reading.

- M'énard R, Yang Y and Rochon Y. 2009. Convergence and stability of estimated error variances derived from assimilation residuals in observation space. In: *Proceedings of ECMWF Workshop on diagnostics of data assimilation system performance*, 15-17 June 2009.
- Livings DM, Dance SL and Nichols NK. 2008. Unbiased Ensemble Square Root Filters, *Physica D: Nonlinear Phenomena*, 237, 1021-1028.
- Smith PJ, Lawless AS and Nichols NK. 2017. Treating sample covariances for use in strongly coupled atmosphere-ocean data assimilation, *Geophysical Research Letters*, 44, <http://onlinelibrary.wiley.com/doi/10.1002/2017GL075534/full>
- Stewart LM, Dance, SL and Nichols NK. 2008. Correlated observation errors in data assimilation. *Int J for Numer Methods in Fluids*, 56:1521–1527.
- Stewart LM, Cameron J, Dance SL, English S, Eyre JR, Nichols NK. 2009. Observation error correlations in IASI radiance data. University of Reading. Dept of Mathematics & Statistics, Mathematics Report 1/2009.
- Stewart LM. 2010. Correlated observation errors in data assimilation. PhD thesis, Dept of Mathematics & Statistics, University of Reading.
- Stewart LM, Dance SL, Nichols NK. 2013. Data assimilation with correlated observation errors: experiments with a 1-D shallow water model *Tellus A*, 65, 2013.
- Stewart LM, Dance SL, Nichols NK, Eyre JR, Cameron J. 2014. Estimating interchannel observation error correlations for IASI radiance data in the Met Office system. *QJ Royal Meteor Soc*, 140:1236-1244.
- Tabcart J. 2016. On the variational data assimilation problem with non-diagonal observation weighting matrices. MRes thesis, Dept of Mathematics & Statistics, University of Reading.

- Tabeart JM, Dance SL, Haben SA, Lawless AS, Nichols NK and Waller JA. 2018. The conditioning of least squares problems in variational data assimilation, *Numer Linear Algebr Appl*, 2018;e2165 (pp. 22).
- Tabeart JM, Dance SL, Haben SA, Lawless AS, Nichols, NK and Waller JA. 2018b. Improving the condition number of estimated covariance matrices, MPECDT Jamboree 2018. Imperial College, poster presentation.
- Waller JA. 2013, Using observations at different spatial scales in data assimilation for environmental prediction, PhD thesis, Dept of Mathematics & Statistics, University of Reading.
- Waller JA, Dance SL, Lawless AS, Nichols NK and Eyre JR. 2014a. Representativity error for temperature and humidity using the Met Office high resolution model. *QJ Royal Meteor Soc*, 140:1189-1197
- Waller JA, Dance SL, Lawless AS and Nichols NK. 2014b. Estimating correlated observation errors with an ensemble transform Kalman filter. *Tellus A*, 66, 23294 (15 pp) .
- Waller JA, Dance SL and Nichols NK. 2016a. Theoretical insight into diagnosing observation error correlations using background and analysis innovation statistics, *QJ Royal Meteor Soc*, 142 (694), pp. 418-431.
- Waller JA, Simonin D, Dance SL, Nichols NK and Ballard SP. 2016b. Diagnosing observation error correlations for Doppler radar radial winds in the Met Office UKV model using observation-minus-background and observation-minus-analysis statistics. *Mon Wea Rev*, 144, 3533–3551. doi: 10.1175/MWR-D-15-0340.1.

- Waller JA, Ballard SP, Dance SL, Kelly G, Nichols NK and Simonin D. 2016c. Diagnosing horizontal and inter-channel observation error correlations for SEVIRI observations using observation-minus-background and observation-minus-analysis statistics. *Remote Sensing*, 8, 581 (14pp).
- Waller JA, Dance SL and Nichols NK. 2017. On diagnosing observation error statistics with local ensemble data assimilation, *QJ Royal Meteor Soc*, 143, 2677 – 2686. doi: 10.1002/qj.3117.
- Waller JA, Dance SL, Lawless AS, Nichols NK and Eyre JR. 2014a. Representativity error for temperature and humidity using the Met Office high resolution model. *QJ Royal Meteor Soc*, 140:1189-1197
- Wattrelot E, Montmerle T and Guerrero CG. 2012. Evolution of the assimilation of radar data in the AROME model at convective scale. In Proceedings of the 7<sup>th</sup> European Conference on Radar in Meteorology and Hydrology.
- Weston PP, Bell W, and Eyre JR. 2014. Accounting for correlated error in the assimilation of high-resolution sounder data. *QJ Royal Meteor Soc*, doi: 10.1002/qj.2306